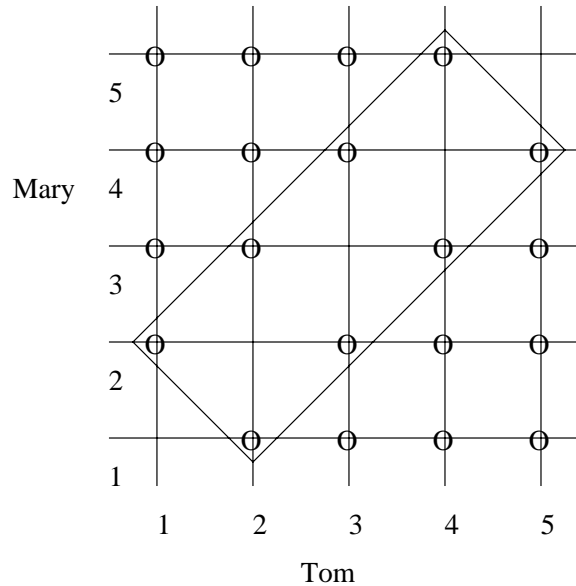


**Problem Set 3: Solutions**  
**Due: March 1, 2006**

1. The problem did not explicitly state that two cars cannot share a parking space, but it was expected that you would assume this when doing the required counting.

The figure below depicts the full outcome space for the case of  $N = 5$ . The 8 outcomes in the box (out of the total of 20 outcomes) are those for which Mary and Tom are parked adjacently.



Extending this idea to a parking lot with  $N$  spaces, the desired probability is given by

$$\begin{aligned}
 \mathbf{P}(\text{parked adjacently}) &= \frac{\text{number of outcomes with adjacent parking}}{\text{total number of outcomes}} \\
 &= \frac{2(N-1)}{N^2 - N} \\
 &= \frac{2}{N}.
 \end{aligned}$$

2. (a) There are nine equally-likely ordered pairs  $(i, j)$ ,  $i \in \{1, 2, 3\}$ ,  $j \in \{1, 2, 3\}$ . By looking at the five possible sums and their frequencies, we obtain

$$p_X(k) = \begin{cases} 1/9, & k = 1; \\ 2/9, & k = 2; \\ 3/9, & k = 3; \\ 2/9, & k = 4; \\ 1/9, & k = 5; \\ 0, & \text{otherwise.} \end{cases}$$

- (b) The fair price is  $\mathbf{E}[5X]$  because then the net expected result is  $\mathbf{E}[5X - a] = 0$ .

$$\mathbf{E}[5X] = \frac{1}{9} \cdot 5 + \frac{2}{9} \cdot 10 + \frac{3}{9} \cdot 15 + \frac{2}{9} \cdot 20 + \frac{1}{9} \cdot 25 = 15$$

(c) The possible values for  $X$  are changed, but the probabilities are unchanged:

$$p_X(k) = \begin{cases} 1/9, & k = 1; \\ 2/9, & k = 4; \\ 3/9, & k = 9; \\ 2/9, & k = 16; \\ 1/9, & k = 25; \\ 0, & \text{otherwise.} \end{cases}$$

$$\mathbf{E}[5X] = \frac{1}{9} \cdot 5 + \frac{2}{9} \cdot 20 + \frac{3}{9} \cdot 45 + \frac{2}{9} \cdot 80 + \frac{1}{9} \cdot 125 = \frac{155}{3}$$

3. Denote the die rolls by  $W$  and  $Z$ . The sixteen equally-likely  $(W, Z)$  ordered pairs are depicted below, where the label in each cell is the  $(X, Y)$  pair.

	$Z = 1$	$Z = 2$	$Z = 3$	$Z = 4$
$W = 1$	(0,1)	(1,1)	(2,1)	(3,1)
$W = 2$	(1,1)	(1,2)	(2,2)	(3,2)
$W = 3$	(2,1)	(2,2)	(2,3)	(3,3)
$W = 4$	(3,1)	(3,2)	(3,3)	(3,4)

(a) From the table, we can read off the PMFs

$$p_X(k) = \begin{cases} 1/16, & k = 0; \\ 3/16, & k = 1; \\ 5/16, & k = 2; \\ 7/16, & k = 3; \\ 0, & \text{otherwise;} \end{cases} \quad \text{and} \quad p_Y(k) = \begin{cases} 7/16, & k = 1; \\ 5/16, & k = 2; \\ 3/16, & k = 3; \\ 1/16, & k = 4; \\ 0, & \text{otherwise,} \end{cases}$$

and thus compute the expectations

$$\mathbf{E}[X] = \frac{1}{16} \cdot 0 + \frac{3}{16} \cdot 1 + \frac{5}{16} \cdot 2 + \frac{7}{16} \cdot 3 = \frac{17}{8}$$

and

$$\mathbf{E}[Y] = \frac{7}{16} \cdot 1 + \frac{5}{16} \cdot 2 + \frac{3}{16} \cdot 3 + \frac{1}{16} \cdot 4 = \frac{15}{8}.$$

We get by linearity of the expectation that  $\mathbf{E}[X - Y] = \mathbf{E}[X] - \mathbf{E}[Y] = \frac{1}{4}$ .

(b) Using the PMFs in part (a), we can compute

$$\mathbf{E}[X^2] = \frac{1}{16} \cdot 0^2 + \frac{3}{16} \cdot 1^2 + \frac{5}{16} \cdot 2^2 + \frac{7}{16} \cdot 3^2 = \frac{43}{8}$$

and

$$\mathbf{E}[Y^2] = \frac{7}{16} \cdot 1^2 + \frac{5}{16} \cdot 2^2 + \frac{3}{16} \cdot 3^2 + \frac{1}{16} \cdot 4^2 = 30.$$

Thus,  $\text{var}(X) = \mathbf{E}[X^2] - (\mathbf{E}[X])^2 = \frac{55}{64}$  and  $\text{var}(Y) = \mathbf{E}[Y^2] - (\mathbf{E}[Y])^2 = \frac{1695}{64}$ .

MASSACHUSETTS INSTITUTE OF TECHNOLOGY  
Department of Electrical Engineering & Computer Science  
**6.041/6.431: Probabilistic Systems Analysis**  
(Spring 2006)

---

Since  $X$  and  $Y$  are *not* independent, the variance of  $X$  and  $Y$  is not any simple combination of previous results. Instead, let  $Z = X - Y$  and find the PMF of  $Z$  as

$$p_Z(k) = \begin{cases} 4/16, & k = -1; \\ 6/16, & k = 0; \\ 4/16, & k = 1; \\ 2/16, & k = 2; \\ 0, & \text{otherwise.} \end{cases}$$

Now

$$\mathbf{E}[Z^2] = \frac{4}{16} \cdot (-1)^2 + \frac{6}{16} \cdot 0^2 + \frac{4}{16} \cdot 1^2 + \frac{2}{16} \cdot 2^2 = 1,$$

and  $\text{var}(Z) = \mathbf{E}[Z^2] - (\mathbf{E}[Z])^2 = 1 - (1/4)^2 = \frac{15}{16}$ . ( $\mathbf{E}[Z]$  was computed in part (a) and can also be double-checked with the PMF above.)

We will use the formula

$$\text{var}(Y) = \mathbf{E}[Y^2] - (\mathbf{E}[Y])^2$$

for the variance of a random variable  $Y$ . Let  $Y = (X - \hat{x})$ . Then

$$e(\hat{x}) = \mathbf{E}[(X - \hat{x})^2] = \text{var}(X - \hat{x}) + (\mathbf{E}[X - \hat{x}])^2 = \text{var}(X) + (\mathbf{E}[X] - \hat{x})^2,$$

where the last equality follows from the fact that shifting a random variable by a constant (in this case  $\hat{x}$ ) does not change its variance. Since the first term is not dependent on  $\hat{x}$  and the second is always nonnegative, we see that this expression is minimized when  $\mathbf{E}[X] - \hat{x} = 0$ . This is equivalent to the desired result of  $\hat{x} = \mathbf{E}[X]$ .

4. (a) From the joint PMF, there are six  $(x, y)$  coordinate pairs with nonzero probabilities of occurring. These pairs are  $(1, 1)$ ,  $(1, 3)$ ,  $(2, 1)$ ,  $(2, 3)$ ,  $(4, 1)$ , and  $(4, 3)$ . The probability of a pair is proportional to the product of the  $x$  and  $y$  coordinate of the pair. Because the probability of the entire sample space must equal 1, we have:

$$(1 \cdot 1)c + (1 \cdot 3)c + (2 \cdot 1)c + (2 \cdot 3)c + (4 \cdot 1)c + (4 \cdot 3)c = 1.$$

Solving for  $c$ , we get  $c = \boxed{\frac{1}{28}}$

- (b) There are three sample points for which  $Y < X$ .

$$\mathbf{P}(Y < X) = \mathbf{P}(\{(2, 1)\}) + \mathbf{P}(\{(4, 1)\}) + \mathbf{P}(\{(4, 3)\}) = \frac{2 \cdot 1}{28} + \frac{4 \cdot 1}{28} + \frac{4 \cdot 3}{28} = \boxed{\frac{18}{28}}$$

- (c) There are two sample points for which  $Y > X$ .

$$\mathbf{P}(Y > X) = \mathbf{P}(\{(1, 3)\}) + \mathbf{P}(\{(2, 3)\}) = \frac{1 \cdot 3}{28} + \frac{2 \cdot 3}{28} = \boxed{\frac{9}{28}}$$

- (d) There is only one sample point for which  $Y = X$ .

$$\mathbf{P}(Y = X) = \mathbf{P}(\{(1, 1)\}) = \frac{1 \cdot 1}{28} = \boxed{\frac{1}{28}}$$

Notice that, using the above two parts:

$$\mathbf{P}(Y < X) + \mathbf{P}(Y > X) + \mathbf{P}(Y = X) = \frac{18}{28} + \frac{9}{28} + \frac{1}{28} = 1$$

as expected.

- (e) There are three sample points for which  $y = 3$ .

$$\mathbf{P}(Y = 3) = \mathbf{P}(\{(1, 3)\}) + \mathbf{P}(\{(2, 3)\}) + \mathbf{P}(\{(4, 3)\}) = \frac{3}{28} + \frac{6}{28} + \frac{12}{28} = \boxed{\frac{21}{28}}$$

- (f) In general, for two discrete random variables  $X$  and  $Y$  for which a joint PMF is defined, we have

$$p_X(x) = \sum_{y=-\infty}^{\infty} p_{X,Y}(x, y) \quad \text{and} \quad p_Y(y) = \sum_{x=-\infty}^{\infty} p_{X,Y}(x, y).$$

In this problem the number of possible  $(X, Y)$  pairs is quite small, so we can determine the marginal PMFs by enumeration. For example,

$$p_X(2) = \mathbf{P}(\{(2, 1)\}) + \mathbf{P}(\{(2, 3)\}) = \frac{8}{28}.$$

Overall, we get:

$$p_X(x) = \begin{cases} 4/28, & x = 1; \\ 8/28, & x = 2; \\ 16/28, & x = 4; \\ 0, & \text{otherwise} \end{cases} = \begin{cases} 1/7, & x = 1; \\ 2/7, & x = 2; \\ 4/7, & x = 4; \\ 0, & \text{otherwise} \end{cases}$$

and

$$p_Y(y) = \begin{cases} 7/28, & y = 1; \\ 21/28, & y = 3; \\ 0, & \text{otherwise} \end{cases} = \begin{cases} 1/4, & y = 1; \\ 3/4, & y = 3; \\ 0, & \text{otherwise.} \end{cases}$$

- (g) In general, the expected value of any discrete random variable  $X$  is given by

$$\mathbf{E}[X] = \sum_{x=-\infty}^{\infty} xp_X(x).$$

For this problem,

$$\mathbf{E}[X] = 1 \cdot \frac{1}{7} + 2 \cdot \frac{2}{7} + 4 \cdot \frac{4}{7} = \boxed{3}$$

and

$$\mathbf{E}[Y] = 1 \cdot \frac{1}{4} + 3 \cdot \frac{3}{4} = \boxed{\frac{5}{2}}$$

- (h) The variance of a random variable  $X$  can be computed as  $\mathbf{E}[X^2] - \mathbf{E}[X]^2$  or as  $\mathbf{E}[(X - \mathbf{E}[X])^2]$ . Here we use the second approach.

$$\text{var}(X) = (1 - 3)^2 \cdot \frac{1}{7} + (2 - 3)^2 \cdot \frac{2}{7} + (4 - 3)^2 \cdot \frac{4}{7} = \boxed{\frac{10}{7}}$$

$$\text{var}(Y) = \left(1 - \frac{5}{2}\right)^2 \frac{1}{4} + \left(3 - \frac{5}{2}\right)^2 \frac{3}{4} = \frac{9}{16} + \frac{1}{16} = \boxed{\frac{5}{8}}$$

G1<sup>†</sup>. Starting with the hint, we have

$$\mathbf{E}[(\alpha X + Y)^2] \geq 0,$$

which can be expanded to

$$\alpha^2 \mathbf{E}[X^2] + 2\alpha \mathbf{E}[XY] + \mathbf{E}[Y^2] \geq 0.$$

The lack of real solutions  $\alpha$  to

$$\alpha^2 \mathbf{E}[X^2] + 2\alpha \mathbf{E}[XY] + \mathbf{E}[Y^2] = \beta$$

for any  $\beta < 0$  implies that the discriminant of the above quadratic,  $(2\mathbf{E}[XY])^2 - 4\mathbf{E}[X^2]\mathbf{E}[Y^2]$ , must be nonpositive. Rearranging

$$(2\mathbf{E}[XY])^2 - 4\mathbf{E}[X^2]\mathbf{E}[Y^2] \leq 0$$

gives the desired result.